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3 **Predicting foraging wading bird populations in Everglades National Park**  
4 **from seasonal hydrologic statistics under different management scenarios**  
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## Abstract

The ability to map relationships between ecological outcomes and hydrologic conditions in Everglades National Park (ENP) is a key building block for the restoration program, a primary goal of which is to improve conditions for wading birds. This paper reports on a model linking wading bird foraging numbers to hydrologic conditions in ENP. We demonstrate that seasonal hydrologic statistics derived from a single water level recorder are well-correlated with water depths throughout most areas of ENP, and are effective as predictors of wading bird numbers when using a nonlinear hierarchical Bayesian model to estimate the conditional distribution of bird populations given the seasonal statistics at the index location. Model parameters are estimated using a Markov Chain Monte Carlo procedure. Parameter and model uncertainty are assessed as a byproduct of the estimation process. Water depths at the beginning of the nesting season, the dry season recession rate, and the numbers of reversals in the recession are identified as significant predictors, consistent with the hydrologic conditions considered important in the production and concentration of prey organisms in this system. Long-term hydrologic records at the index location allow for a retrospective analysis (1952-2006) of foraging bird numbers showing low frequency oscillations in response to decadal fluctuations in hydroclimatic conditions. Simulations of water levels at the index location used in the Bayesian model under alternative water management scenarios produce variable numbers of foraging birds and the results provide criteria for linking management schemes to seasonal rainfall forecasts.

Keywords: Everglades National Park, Wading Bird, Hydrology, Bayesian Model

## 1. Introduction

Many of the adverse impacts on ecosystem function in the Everglades National Park (ENP; **Figure 1**) are caused directly or indirectly by altered regional hydrology [Bancroft *et al.*, 2002; Craighead, 1971; Davis, 1943; Gleason, 1984; Loveless, 1959; Powell *et al.*, 1989]. However, it is difficult to identify cause-and-effect relations between altered hydrologic patterns and ecosystem function in ENP because of the many factors involved, and because a long-term data record on each of the putative factors is needed to assess the relationships. Consequently, modeling studies [Cline *et al.*, 2006; Gaines, 2000; Wetzel, 2001] have been undertaken to create synthetic records that could represent the hydro-ecology of the region. Since the relationships in such models are largely prescribed or empirically estimated from relatively short records, it is difficult to rely solely on these models as tools for understanding or predicting the ecological outcomes in ENP consequent to hydroclimatic variations. A few authors [Bancroft *et al.*, 2002; Russell *et al.*, 2002] have tried to directly use data on a few species and hydrologic or climatic indicators to infer such relationships that could be used for subsequent management of ENP. This paper contributes to this literature.

A goal of ENP restoration project is to ensure that the ecological health of ENP improves as a direct result of management activities. Achieving hydrologic targets through the proper timing and amount of releases from control structures is a first step in the management process. Significant climate and weather variations in the region influence the ability to make releases and also determine the ecological outcomes. A predictive model for ecological outcomes given anticipated climate conditions and proposed releases is a basic building block for an adaptive management process. The development of such a model with a capacity for uncertainty analysis is the goal of the larger research project to which this paper contributes.

Seasonal water depths in ENP depend on managed surface water releases from control structures and on direct rainfall. On an annual basis and for ENP as a whole, direct rainfall is the dominant component of the two water sources. However, the importance of the discharges through the control structures increases during the dry season. These discharges are concentrated along the northern boundary of ENP and they have their largest influence in the region directly downstream called Shark Slough. The ecological responses to these discharges must be quantified in order to effectively manage the Park and to guide the restoration effort. However, given the large influence of rainfall on the system, the ecological responses to the managed discharges must be placed within the context of larger-scale climatic factors. Here we link the variations in the foraging populations of two wading bird species – a fundamental aspect of Everglades' ecology – to hydrologic conditions in the National Park that result from rainfall and managed releases.

Foraging patterns of wading birds have been a key issue in the Comprehensive Everglades Restoration Plan (CERP), and monitoring of wading bird nesting success is a coordinated effort between many agencies in Florida. Although other factors (e.g. prey conditions, migrations from remote areas) may influence the foraging patterns of wading birds in ENP [Cristol and Switzer, 1999; Gawlik, 2002; Houtman and Dill, 1998; Krebs and Cowie, 1976; Lima and Dill, 1990; Safina and Burger, 1985], it is likely that hydrologic conditions are a major driver that dominate the underlying population dynamics in the Park [Cezilly *et al.*, 1995; Frederick and Collopy, 1989; Gawlik, 2002; Powell, 1987; Russell *et al.*, 2002; Smith, 1995; Spalding *et al.*, 1993; Strong *et al.*, 1997]. Previous studies [Russell *et al.*, 2002] have shown an inverse relationship between the number and degree of dry season disruptions (short term reversals in the recession of surface water) and wading bird foraging patterns.

Our study seeks to quantify the response of wading bird foraging numbers to selected seasonal statistics of the daily water level. The long term goal is to improve wading bird outcomes in ENP through the use of climate forecasts and adaptive water release policies from control structures such that a better water level trajectory results for the upcoming season from the perspective of wading bird foraging success. This entails climate and water level forecasts in ENP for each candidate release policy and the connection of the resulting water level forecasts to potential wading bird populations. This paper presents one critical piece of this exercise – the ability to predict seasonal wading bird foraging numbers from water level statistics at a single, centrally-located gauge (P33) located below the control structures. We show that water levels through much of ENP, especially those in regions important for wading bird foraging, are highly correlated with water levels at P33. The relatively long record (>50 yrs) at this gauge also provides a basis for linking wading bird foraging patterns in ENP to decadal-scale climate variability.

This paper develops a Hierarchical Bayesian Model that relates the population of two dominant wading birds, the Great Egret (*Casmerodius albus*) and White Ibis (*Eudocimus albus*), to key seasonal statistics of the water levels at P33. Water levels at P33 and potential wading bird foraging populations are then predicted, again using a Bayesian scheme, based on rainfall and inflow volumes under different management scenarios for the Park. We take a Bayesian approach so that model and parameter uncertainty can be quantified. Probabilistic, seasonal climate forecasts will eventually be used for decision analysis regarding managed releases, and since these decisions must also incorporate uncertainty information to ensure a specified reliability in outcomes, a Bayesian approach for the entire process is attractive to properly assess the propagation of this uncertainty in modeling and analysis.

## 2. Study Area and Data

The greater Everglades ecosystem extends from the southern edge of Lake Okeechobee to Florida Bay. The southern terminus of the system has been preserved as ENP, one of the most widely recognized wetlands in the world. The historic ecosystem was once characterized by large expanses of shallow, slowly moving surface water (called sheetflow) with seasonally fluctuating water levels controlled primarily by rainfall and runoff from Lake Okeechobee. Hydropatterns over much of this region have been altered through various forms of land use change and management practices related to regional water supply and flood control. The once contiguous wetlands have been compartmentalized by a series of canals and levees that, along with a network of pumps, weirs, and gated culverts, are the primary means by which water levels are controlled and deliveries made to the population centers and natural areas. As a result of these modifications, sheetflow patterns in the system have been altered, causing deviations from historic water level fluctuations and an overall decline of ecosystem function within ENP. Most notably, wading bird populations are estimated to have declined by 90% relative to their historic levels [Ogden, 1994]. ENP has a sub-tropical climate with a distinct wet season in the summer and a dry season in the winter. Almost, 75% of the annual precipitation falls during May-October with monthly precipitation amounts ranging between 0.0 and 20.0 inch. Overland sheetflow from northern contributing zones into ENP is a fraction of the historical amounts. Ecologic restoration of ENP brings to the fore the challenge of how to deliver the right amount of water to the Park at the right times to the right locations.

The primary data considered are historical rainfall records, water stages and the foraging abundance of *Great Egret* and *White Ibis*. Water level data are taken from the P33 gauge which is shown below to be directly related to overall hydrologic conditions in the Park, and for which

long-term reliable data are readily available from 1952 to 2007. Seasonal variation in water stage at P33 is illustrated in **Figure 2(a)**.

Wading bird foraging data from the Systematic Reconnaissance Survey Flights (SRF) [Norton-Griffiths, 1978; Russell *et al.*, 2002] were used. The SRF involve flying at a fixed altitude and speed across a study area on a predetermined transect while observers count animals in a strip of land on either side of the aircraft. The SRF surveys were initiated to give South Florida's operational resource managers a tool to assess wildlife populations. Past work to explore the effects of hydrology on SRF-derived wading bird distribution and abundance data is documented in Russell *et al.*, [2002], Bancroft *et al.*, [1992, 2002], and Porter and Smith [1984]. The SRF wading bird survey was initiated in 1985. The data are updated every month from December through May, the dry season in south Florida, and once in August. [Russell *et al.*, 2001] provide details of the SRF wading bird surveys. In the present work the wading bird data were derived by aggregation from any grid cell in which each species was ever recorded by SRF from 1985 to 2006. Seasonal variations in *Great Egret* and *White Ibis* populations are shown in **Figure 2(b-c)**.

## **3.0 Preliminary Analyses**

### **3.1 Relationship between P33 stage and water depths in ENP**

During the dry season, large numbers of wading birds forage within ENP and tend to concentrate along Shark Slough. Aerial survey records (1985-present) show the numbers of wading birds foraging in the Slough and throughout ENP fluctuate on an annual basis. These fluctuations have been linked to water depths at the beginning of the dry season and the subsequent recession rates by [Russell *et al.*, 2002]. However, the hydrologic data used by

[Russell et al., 2002] were derived from qualitative aerial observations and not based on actual gauge data. This is because water level recorders have been scattered non-uniformly in the Park and operational for different time periods. Synoptic water depths for the entire Park were therefore not available for the period of record matching the bird surveys. Recent installations of automated gauge stations have increased the spatial coverage of water level measurements and allowed for interpolation between these stations. Updated, grid-based (400 m<sup>2</sup>) topography data [Desmond, 2003] facilitates the conversion of the interpolated water levels to water depths for most of ENP.

Water level fluctuations across ENP are highly correlated. A linear correlation map between daily water stages from January to April 2006 at P33 and daily water depths derived from interpolation of gauge data throughout ENP is illustrated through the colored contour maps in **Figure 3**. A high correlation is observed between stages at P33 and water depths throughout ENP over this period. The correlations are highest along the longitudinal axis of Shark Slough and in areas of similar elevations during both the dry season and wet season. Control structures discharge water across the northern boundary of ENP directly into the Slough and have less effect on water levels in adjacent areas of higher elevation, or in areas close to the border canals. Thus, there is some decrease in correlation between stages at P33 and water depths outside of the Slough, such as in some of the marl prairie regions. For the period from 2000-2007 we found an average  $r^2$  correlation coefficient of 0.75 (ranging from 0.66 to 0.95) between January-April daily water depths at P33 and water levels at seven principal gauges (NE2, NP201, EVER6, CP, P36, NP46, CR2) in ENP, each representing one of the landscape classifications shown in Figure 3. With this information we consider water levels at P33 a useful indicator of water depths



throughout most of ENP, particularly in areas of Shark Slough where wading bird foraging is concentrated.

Observations of foraging *Great Egret* and *White Ibis* in May of each year (1985-2006) and are also displayed in Figure 3. High wading bird counts can be noted in those regions of ENP that show the highest correlation with P33 water levels. As a consequence, water levels at P33 may be useful to infer the suitability of hydrologic conditions for wading birds throughout the Park.

### 3.2 Identifying Suitable Predictor Variables from the P33 data

Bird foraging patterns are affected by many different factors such as prey availability [Gawlik, 2002; Krebs and Cowie, 1976; Lima and Dill, 1990; Safina and Burger, 1985] and the threat of predation [Cristol and Switzer, 1999; Houtman and Dill, 1998; Lima and Dill, 1990]. Hydrologic conditions are also known to play a role [Powell 1987]. Changes in the historical pattern of water level fluctuation are considered to be a significant factor that determines foraging patterns and the abundance of bird populations in ENP [Cezilly *et al.*, 1995; Frederick and Collopy, 1989; Gawlik, 2002; Powell, 1987; Smith, 1995; Spalding *et al.*, 1993; Russell *et al.*, 2002; Strong *et al.*, 1997].

It is important to recognize that hydrologic conditions may be related to other factors influencing wading bird foraging [Gawlik, 2002]. For example, the density and distribution of vegetation in ENP is an example of a habitat factor related to hydrologic conditions that may affect foraging patterns. Marsh vegetation types and density in the Everglades are known to change with alterations to hydroperiod [Ross *et al.*, 2003; Armentano *et al.*, 2006; Saunders *et al.*, 2006], and these changes in vegetation distribution/density may affect the habitat quality for foraging wading birds. The populations of wading birds in areas outside ENP, (e.g. the Water

Conservation Areas) and the migrations of birds to and from these areas may also be a determinant of the foraging patterns observed during the SRF flights. The impact of these migrations on the SRF observations is not well known and was not included in our model. Since little information on the impacts of factors such as vegetation types or migration on foraging patterns is available, we focus on hydrologic statistics of within-season variation in P33 stage as potential predictors of foraging numbers for the two species of interest.

Everglades wading birds nest in the December – May dry season, and successful foraging during this period can be considered a prerequisite for successful nesting [Russell *et al.*, 2002; Frederick and Ogden, 2003]. Foraging success in the dry season is dependent on water depths being low enough to allow standing, and on appropriate surface water recession rates which concentrate prey in the low-lying areas [Kushlan, 1986; Frederick and Collopy, 1989; Gawlik, 2002]. Disruptions to the dry season recession caused by winter rainfall events or from managed water releases tend to reduce foraging success because of the resulting prey dispersal [Frederick and Ogden, 2003]. Too rapid recession rates or too low water levels in December may shorten the nesting season, although it is important to consider that the impacts of these and other hydrologic factors on foraging/nesting are often species-dependent [Frederick and Spalding, 1994]. For this initial analysis, the number of *Great Egret* and *White Ibis* in May at the end of the dry season are selected as the predictands reflecting aggregate hydrologic variables calculated from January - April. These two species were chosen because their white color makes them easy to identify in the SRF flights, and these data are therefore less subject to error. As in Russell *et al.* (2002) bird counts in May were chosen because late-dry season foraging numbers can be considered an indicator of nesting initiated earlier in the season. Similar models could be developed using data collected during any dry season month and for other species of wading

birds. Models which incorporate within-season variability in foraging and nesting patterns with changes in the regional and local-scale hydro patterns will improve understanding of the specific relationships between hydrology, foraging, and the success of individual colonies [Bancroft *et al.*, 1994]. The objective of this paper represents a step in this direction by demonstrating the development of a robust analytical framework capable of linking ecological outcomes to hydrologic indicators in a manner that is useful to managers and restoration planners.

Now we consider the seasonal statistics of the daily water levels recorded at P33 that may be useful predictors of the May bird count. The approach followed is generally similar to that used by [Russell *et al.*, 2002], except that the predictors are derived from actual daily water level data from a single location, P33, instead of using the gridded qualitative aerial observations of water level at a monthly scale.

The procedure used to develop seasonal water stage statistics as predictors from the daily water stage data at P33 is illustrated in **Figure 4**. First, consider a linear decline of stage with time during the middle of the dry season (January through April) representing the seasonal recession of the water table. For this recession, a linear regression of stage versus time into the season provides the intercept as an estimate of the mean of initial water level, and the slope as an estimate of the recession rate. Once this recession behavior is estimated, a disruption can be defined as a positive residual from the linear regression line. For each season, we can then compute the number of disruptions, a standard deviation of disruption, a maximum consecutive disruption, the average water stage, the intercept and the recession slope at P33 as potential predictors.

A stepwise regression procedure together with exploratory data analysis was used next to screen these potential predictors for each bird count. Finally, the initial water level, the averaged

water level and the number of disruptions from January through April were selected as predictors. A smooth surface fit using a Cross validated Thin-plate smoothing spline [Wahba, 1990] of the birds as a function of initial water stage and disruption is shown in **Figure 5**. Note that the bird counts are positively correlated with initial stage and average stage, and are negatively correlated with the number of disruptions as can be expected. As may also be expected, the initial and average stage for the season are correlated. The correlations across the three predictors and the two predictands, and partial correlations for each prediction are provided in **Table 1**. Based on these partial correlations it is apparent that while the predictors are correlated with each other, they still contribute useful additional predictive information. An examination of Figure 5 illustrates that the relationship of bird count with water level is nonlinear. Bird counts initially increase with average or initial water level, but decrease or level out for high stage. Consequently, a quadratic relationship between the predictand and these predictors may be appropriate [Russell *et al.*, 2002].

A K-means [Spath, 1985] cluster analysis was applied to the standardized time series of the three predictors and population counts for each bird species. The intention was to see how these attributes group together. Box plots of the number of birds and the predictors for the three clusters identified are shown in **Figure 6**. The first cluster corresponds to medium water stage and relatively low disruption. The median bird populations are the highest for this cluster. The White Ibis population has high variation for this cluster, but is still generally higher than for the other clusters. Cluster 2 has the highest water levels and a medium level of disruption. It maps to a middle category of median bird population, but with high variation in the Great Egret bird count. The third cluster corresponds to the lowest water levels and the highest disruption

frequency and maps on to the lowest bird counts for both species. The nonlinearity of the relationship between the predictors and the predictand is further illustrated by this analysis.

#### 4.0 A Predictive Model for Foraging Birds Populations

A hierarchical Bayesian model is developed for the seasonal prediction of seasonal populations of *Great Egret* and *White Ibis* using the selected hydrologic variables, specifically the initial water stage, the average water stage and the number of disruptions. The model considers that the population of each bird species follows a log normal distribution with time varying mean  $\mu(t)$  and a constant variance  $\sigma$ . A quadratic model for the mean  $\mu_j(t)$  for bird species  $j$  in terms of each predictor is then formulated as follows:

$$(1) \quad Z_j(t) \sim LN(\mu_j(t), \sigma_j)$$

$$(2)$$

$$\mu_j(t) = \beta_{1j} + \beta_{2j} \cdot IWS(t) + \beta_{3j} \cdot IWS(t)^2 + \beta_{4j} \cdot AWS(t) + \beta_{5j} \cdot AWS(t)^2 + \beta_{6j} \cdot DIS(t) + \beta_{7j} \cdot DIS(t)^2$$

Where  $Z_j(t)$  represents the bird count for species  $j$  for season  $t$ , with mean  $\mu_j(t)$  and standard deviation  $\sigma_j$  (the variance is taken to be a constant over time after some initial experimentation); *IWS* is initial water stage, *AWS* is averaged water stage and *DIS* is number of disruptions at P33 from January to April.

Each of the parameters in the model above is considered as a random variable with a given probability distribution. A schematic of the model is presented in **Figure 7**. We note that the correlation of the observed abundance between the two bird species is 0.85, and hence it may be useful to consider pooling these data in a regression conditioned on the same predictors. The

hierarchical Bayesian model used here provides an objective way to choose the degree of pooling. First we apply the regression to foraging bird population data standardized by dividing the raw series by the mean annual bird population for the corresponding series. Next, we assume that the regression coefficients for each bird species for each predictor come from a common distribution with a common mean and variance. If this variance is small then effectively we have a pooled regression. Conversely, if this variance is large, then independent regressions for the two species result. If the number of bird species introduced into the model increases, then this procedure would still apply and would lead to a better estimate of the mean and variance of the common regression coefficient.

The hierarchical regression models are solved in a Bayesian framework. Non-informative priors are assumed for each of the parameters and hyperparameters (e.g.,  $\beta_{ij}$ ,  $\mu_{\beta_{ij}}$ ,  $\mu_{\mu_{\beta_{ij}}}$ ,  $\sigma_j$ ) and their optimal values are selected through a maximization of the posterior likelihood of observing the data. A Markov Chain Monte Carlo (MCMC) procedure is used. In particular, the Gibbs sampling approach to MCMC [Gilks *et al.*, 1995] has been used in this study. We chose to run three chains simultaneously searching for optimal parameters. The evolution of each chain was monitored to check for convergence to a common value. Selection of the hyper-priors and the appropriateness of the prior distributions and the model structure were judged by the deviance information criterion (DIC) [Berg *et al.*, 2004]. All computations were performed in Matlab using Winbugs (Spiegelhalter *et al.*, 2003). To assess convergence for each parameter the [Gelman and Rubin, 1992] ‘shrink factor’ was computed. This factor compares the variation in the sampled parameter values within and between chains, and it describes how much the increase in the number of iterations improves the estimates. [Gelman and Rubin, 1992] suggest running Gibbs sampler chains until the estimated shrink factors are less than about 1.2 for all parameters.

Winbugs produces scale reduction factors that are very close to 1 for the fixed effects. Values around 1 were obtained after 10,000 iterations.

**Table 2** summarizes key results for each regression coefficient. The posterior mean, standard deviation, and 95% credible interval are derived using MCMC as noted above (the detailed estimates of Hyperparameters are summarized in Appendix). Recall that the initial and average water stages are correlated predictors. Hence, it is not possible to uniquely identify one or the other as a predictor, and given the results in Table 1, both the linear and the quadratic terms in these two variables are retained. For both species, the quadratic term in number of disruptions has a posterior 95% interval that brackets zero. This observation is consistent with the observation from the relationship in Figure 5, where the relationship appears linear. So, one could consider the deletion of this term from the model. The fit of the various models was compared using the DIC as calculated in Winbugs [Spiegelhalter *et al.*, 2003]. The lower the DIC value, the better the model fits the data. The model that includes the quadratic term has a lower DIC score than the model that does not include it, so we retained this term as well. Unlike the case for classical linear regression, where predictors are selected and putative causes are argued for based on a hypothesis test for the significance of particular regression coefficients, in the Bayesian formulation, the intention is to present posterior conditional distributions for the  $Z(t)$  that reflect both the uncertainty in estimating the parameters, and the model uncertainty, i.e., the variation in  $Z(t)$  not explained by the model. By including an additional term in the model and considering the uncertainty in it, we may be understating the standard deviation  $\sigma$  ascribed to variations around the conditional mean, i.e., the residual or noise term. However, if the interest is in a posterior interval for  $Z(t)$ , rather than in the attribution of variance to the noise term or to a predictor, then the inclusion of an additional predictor does not change the results.

The time series of observed wading birds and the values predicted at ENP in May, using Bayesian quadratic regression with the three predictors for 1985–2006 are shown in **Figure 8**. The posterior mean and the 5% to 95% posterior uncertainty bounds are provided. The posterior mean corresponding to a cross validation procedure is also shown. The cross-validation process as implemented here removes seven years from the training set, estimates the regression coefficients with the remaining data, and then predicts the birds for the year that was omitted.

Model predictive ability is quantitatively assessed to judge the degree to which the model simulation matches the actual observations. One can utilize different statistics of efficiency to measure the goodness of fit or prediction skill. [Legates and McCabe, 1999] have critically reviewed many of the principal statistics of efficiency. Nash and Sutcliffe, [1970] defined the Coefficient of Efficiency (CoE) which ranges from minus infinity (poor model) to 1.0 (perfect model). The CoE has been widely used to evaluate the performance of prediction models (e.g., [Wilcox *et al.*, 1990]). For more details regarding goodness of fit measures, see [Legates and McCabe, 1999] and [Willmott *et al.*, 1985]. The statistics used in our study are summarized in **Table 3** and the performance of the model according to these measures is shown in **Table 4**. The *Great Egret* and *White Ibis* posterior means predicted by the model have a correlation of approximately 0.8 with the observed series with a CoE 0.7. The model was also applied to the data from 2000 to 2006, which were originally withheld from the fitting process. No significant change in results is noted. Analysis of residuals (observed –posterior mean) for the model for *Great Egret* and *White Ibis* supported the assumption of zero mean, independently and identically normally distributed errors. Jarque-Bera test for goodness-of-fit to a normal distribution [Judge *et al.*, 1988] and the Shapiro-Wilk parametric hypothesis test of composite



normality [Royston, 1995] were performed, and the hypothesis that error has a normal distribution is not rejected at the 5% and 10% level.

The model developed above is now applied for retrospective prediction of wading bird foraging numbers for the period 1952-1984 for which daily P33 stage data is available but bird data is not. The results are illustrated in **Figure 9**. The long-term trends for P33 average water stage, disruption and rainfall are indicated by Lowess [Cleveland, 1979] smooth lines superimposed on graphs of time series in **Figure 10**. The concurrent decadal to multi-decadal variability in the relevant time series is shown. An interesting feature is the increase in foraging bird population post-1995, even though the seasonally averaged rainfall and P33 stage do not show any marked trends. We note, however, that the disruption time series reveals a decreasing trend over the same time period, reinforcing the earlier assessment that both the stage and the disruption frequency are important indicators for foraging conditions. The trends in foraging numbers in the period before systematic observations are available are consistent with our expectations from the nonlinear model. The period from 1962 to 1970 is marked by relatively low water levels and higher disruption frequency, and translates into lower bird counts. On the other hand the higher bird counts around 1960 and 1980 again correspond to average to higher water levels with a lower disruption frequency. These results rely on the assumption that the relationship between P33 and water levels in the other areas of ENP where wading birds are known to forage remained relatively constant over the complete period of record. Some changes in water management regulation schedules and policies have occurred during this time period. However, the central location of P33 in Shark Slough gives confidence that it has consistently reflected the general hydrologic condition of the areas where wading birds tend to forage. Our goal in presenting these results is to set the stage for a comparative evaluation of different release

policies from the control structures that translate into different AWS, IWS, and DIS values at P33 and therefore, different numbers of foraging bird populations.

## 5. Modeling water levels at P33

A Bayesian approach is developed to relate observed water levels at P33 to a combination of hydrologic predictors representing the water budget parameters at the site. The model considers that over time period  $t$  the water level at P33,  $Z(t)$ , is normally distributed with a time-varying mean  $\mu(t)$  and a constant variance  $\sigma$ . A quadratic model of the mean  $\mu(t)$  in terms of each predictor is then formulated as follows:

$$Z(t) \sim N(\mu(t), \sigma) \quad (3)$$

$$\mu(t) = \beta_1 + \beta_2 \cdot P_{33}(t-1) + \beta_3 \cdot \delta_R(t) \cdot R(t-1) + \beta_4 \cdot \delta_R(t) \cdot R(t-1)^2 + \beta_5 \cdot \delta_I(t) \cdot I(t-1) + \beta_6 \cdot \delta_I(t) \cdot I(t-1)^2 \quad (4)$$

$$\delta_R = \begin{cases} 0 & R \leq 0 \\ 1 & R > 0 \end{cases}, \quad \delta_I = \begin{cases} 0 & I \leq 0 \\ 1 & I > 0 \end{cases} \quad (5)$$

where,  $P_{33}$  is daily average water stage,  $R$  is total daily rainfall and  $I$  is the daily average rate of surface water inflow through upstream control structures into ENP. A delta function is employed to differentiate the zero and the non-zero case for  $R$  and  $I$ . The form of the equation in this case reflects the univariate relationships between the hydrologic parameters known to occur at this site, namely, that mean water levels increases asymptotically with  $I$  and  $R$ , and that  $P_{33}(t)$  is linearly correlated to  $P_{33}(t-1)$ . Non-informative priors were selected and a three chain Gibbs sampling approach to MCMC [Gilks *et al.*, 1995] was employed to solve for the parameters in Matlab using Winbugs [Spiegelhalter *et al.*, 2003]. A Gelman and Rubin [1992] ‘shrink factor’ criteria of 1 in the MCMC was used to establish the parameter values (**Table 5**) and their variation within and between chains. The model coefficient of variation, the coefficient of

efficiency, and the bias show very high fidelity with observed values during separate calibration (1978-1987) and verification (1988-2006) periods.

## **6.0 Impacts of alternative water management policies on water levels and foraging populations**

The impacts of water release policies were investigated by linking predictions of seasonal P33 water levels (and the corresponding AWS, IWS, and DIS) based on different inflow scenarios to the Bayesian model of foraging bird populations. In this scheme P33 water levels were first calculated using (3) & (4) with calibrated parameter values and with variable inflow volumes corresponding to different management scenarios for the period 1986-2000. The “Baseline” case represents P33 and foraging population predictions using measured rainfall and inflows discharging into ENP. Five alternative inflow management scenarios were then tested and are ordered here based on the total amount of water delivered to ENP. Rainfall amounts do not vary among scenarios. The first scenario, “No inflow”, assumes zero releases from the surface water control structures located across the northern boundary of the Park. Surface water releases for the “Minimum inflows” scenario are mandated by federal law (PL 91-282) with the objective of avoiding catastrophic damage to the ecosystem during drought periods [*Light and Dineen* 1994]. The “Rainfall Plan” regulates inflows to ENP as a function of rainfall in the northern contributing basins according to a formula derived by the SFWMD [*Neidrauer and Cooper*, 1989; *Light and Dineen*, 1994]. The Interim Operating Plan “IOP” represents the current management scheme for ENP in place since 2000, while the Combined Structural and Operational Plan “CSOP” scenario is a management scheme accompanying a large-scale restoration project designed to increase and alter the timing of ENP inflows. Inflows into ENP under the IOP and CSOP scenarios were derived from the South Florida Water Management

Model (SFWMM, [http://www.evergladesplan.org/pm/recover/system\\_wide\\_modeling.aspx](http://www.evergladesplan.org/pm/recover/system_wide_modeling.aspx)), a regional-scale hydrology model for south Florida. For Jan-April of each year, the AWS, IWS, and DIS were calculated and used to predict wading bird populations generated in each scenario according to (2).

Aggregate statistics on foraging populations calculated over the full 1986-2000 period of record were found to obscure the differences between scenarios because of the high interannual variability in rainfall and predicted bird populations. We therefore divide the analysis into two time periods representing below-average (1986-1990), and average to above-average (1991-2000) rainfall conditions. During the low rainfall period normalized AWS and IWS are consistently below average for all scenarios (**Table 6**). DIS values are above average for all scenarios during this period, with the exception of the restoration scenario CSOP. The CSOP scenario also produced relatively higher AWS, and together the low DIS this scenario resulted in the highest average numbers of foraging birds during 1986-1990 compared to all other scenarios (**Figure 11**). Student's *t* tests indicate the differences in 1986-1990 populations under these scenarios were not significant due to the short record and high variance caused by the very low rainfall in 1989-1990. However, the results are consistent across scenarios and show that foraging populations during low rainfall periods are predicted to decrease with declining inflows to ENP, with the fewest birds occurring under the “No inflows” and “Minimum Flows” scenarios.

The period 1991-2000 is characterized by average to above-average rainfall, increasing AWS and IWS, and below-average DIS for all scenarios. These conditions resulted in higher numbers of foraging birds compared to the low rainfall period across all scenarios. In general, however, and in contrast to the low rainfall period, the scenarios which delivered relatively less water to ENP (e.g. “No inflows”, “Minimum inflows” and the “Rainfall plan”) produced the

highest numbers of foraging birds. From this analysis, we conclude that during wet periods, high rainfall alone is capable of maintaining AWS and IWS within optimal ranges for wading birds without the addition of managed releases into ENP. DIS values for all scenarios were also lower during high compared to low rainfall conditions and this improved wading bird outcomes. However, the relative decrease in DIS was less apparent under the CSOP scenario. The benefits of higher rainfall to wading bird foraging patterns were therefore reduced in this scenario compared to the others. Anomalous water depths which exceeded the optimal conditions for wading birds during the high rainfall period under both the CSOP and IOP scenarios also reduced the number of birds predicted under these scenarios compared to the others. This is particularly apparent during 1995, when rainfall was ~20% higher (175 cm) than average, and the Great Egret foraging populations under the “No inflow” scenario were 3 and 4 times greater, respectively, than those predicted under IOP or CSOP. Remarkably, the White Ibis populations produced by the “No inflow” scenario were 10-14 times higher than IOP and CSOP during 1995. White Ibis populations in general showed larger responses to management scenarios compared to Great Egrets.

This scenario analysis points to the potential impacts of managed inflows on foraging conditions in ENP. During low rainfall periods, managed releases are necessary to maintain favorable water levels. However, during high rainfall periods these releases may contribute to raising water levels above optimal values. Over the entire 1986-2000 period, the largest average numbers of foraging birds of both species were predicted under the “Minimum inflow” and the “Rainfall Plan” scenarios. However, this is largely because the majority of the period is characterized by high rainfall conditions when the scenarios which restrict inflows performed relatively well. We suggest it is therefore important to consider the performance of these

scenarios in a larger context determined by rainfall timing and amount. Optimizing foraging conditions for wading bird species over longer time periods is likely to require adaptive release schedules which incorporate real-time information on water levels, rainfall, and climatic drivers [e.g. *Kwon et al.*, 2006]. Of the five analyzed here, the management scenario most closely tied to climatic conditions is the “Rainfall Plan”, and while it did not produce the highest numbers of Great Egret or White Ibis during either the low or high rainfall periods, this scenario did produce the highest median number of foraging birds when the species were combined over the full period of record (3200 birds yr<sup>-1</sup>). Similar climate-based approaches to managing inflows into ENP could further improve habitat quality for foraging birds. It is recommended additional analyses be conducted using population data from other species to investigate the broader ecological outcomes of rainfall-driven management plans.

## 7.0 Summary and Discussion

Water depths and their fluctuation in the dry season have an impact on wading bird habitat suitability in ENP. This is well known and many indices have been developed to assess habitat suitability in terms of modeled or observed variables. The work presented here represents a departure from the existing literature that has evolved on this topic. We chose to step back and focus on the predictability of an observed measure of ecological outcomes – estimates of the bird foraging populations – contingent on some identifiable measures of hydrologic conditions in the region. Given the existing literature on the subject, the effort started with modest expectations, and the best and longest available time series for bird populations and a hydrologic time series were used to develop an analysis. A particular challenge in this context was to identify seasonal statistics of daily hydrologic variation that could suitably inform an analysis of aggregate end of season foraging bird populations in a regression framework. After a certain amount of

exploratory work that is presented in a rather condensed form in this paper, a set of predictors that was causally plausible and statistically informative was selected. Initial, nonparametric regression modeling suggested a nonlinear relationship between the response and the predictors. This led to the development of a nonlinear regression model for each bird species in terms of the same set of predictors. Recognition that the populations of the two species are highly correlated with each other suggested that something could be gained by pooling the two regression problems, in terms of characterizing and reducing the uncertainty in the model coefficients. This problem was formulated as a Hierarchical Bayesian regression model and the parameters and hyperparameters of the model were estimated using Markov Chain Monte Carlo methods with a standard approach for checking for convergence and adequacy of model diagnostics. The results of the investigation are presented through posterior conditional densities that include the contribution of parameter uncertainty, and also through cross-validated prediction. The performance statistics are quite impressive for this kind of a model, and provide hope that a simple model such as presented here can be a useful, communicable building block in a strategy that proposes changes in hydrologic conditions through management interventions in a way designed to improve habitat suitability and hence bird outcomes.

Short records and a large number of potential hydroclimatic and other biophysical predictors generally make it difficult to reliably assess the potential response to changed hydrology and the uncertainty associated with models linking ecological outcomes to hydrologic conditions. This challenge is prominent even though several inter-connected spatio-temporal models of hydrology and ecology have been developed specifically for the Park. This situation has been partially addressed in this paper. For instance, knowledge of the initial water level, together with probabilistic climate forecasts, and a proposed water release policy, could be used to estimate

both the probability distributions of average water level in the upcoming season, and the disruption frequency. These could then be jointly used to estimate the probability distribution of birds present at the end of the season using the model developed here. Water release policy refinements could then be pursued in an optimization framework to “deliver” desired population levels with a specified reliability. These policies would include release guidelines that are adaptive within the season, indexed to initial and current water levels, and recent rainfall. Models that map observed and forecast daily rainfall, current stage, and water release from control structures into future stage at P33 would need to be formulated and used as part of a release policy to provide adaptive rules to guide releases such that intra-seasonal hydrologic conditions evolve in a way that leads to improved end-of-season bird outcomes. Work on these aspects is ongoing, and includes verification of the intermediate variables of interest, while connecting the drivers as identified directly to the outcome. Future work will incorporate daily rainfall simulations capturing common or unique aspects of low frequency variation using Wavelet Autoregressive models [ *e.g. Kwon et al., 2006; 2007*].

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## Figure captions

**Figure 1** Everglades National Park is located at the southern tip of the state of Florida, USA. The hydrologic monitoring station P33 is identified by the star near the center of Shark Slough.

**Figure 2** Seasonal trends and variation (box plots) in monthly values of a) water stage at P33, and foraging abundance of b) Great Egret and c) White Ibis from 1985 to 2006. Observations of wading birds are not available in June, July, September, October and November.

**Figure 3** *Great Egret* (a) and *White Ibis* foraging abundance in May 2006, community types, and correlation of water depths to P33 during the dry season. The colored contour map shows the correlation between grid-based water depths in ENP and P33 water levels during the period from January to April 2006. The black circles represent the number of birds in May 2006 and the size of the circle is proportional to the number of birds. The right hand side of the coverage that is not labeled is outside of ENP and subject to different water management practices, resulting in depths with low correlations with P33 and few wading birds. The small area of low correlation on the left hand side indicates a tidal influence not present in other parts of the coverage.

**Figure 4** Three predictors derived from water stage at the P33 station. One is the seasonal average water stage (AWS), the second is the intercept of the best fit line for the decline in stage over the dry season, a measure of the initial water stage (IWS) at the beginning of the foraging season) and the final predictor is the number of disruptions (DIS) to the recession rate with the progression of the dry season.

**Figure 5** Smooth surface fit using a Cross validated Thin-plate smoothing spline [Wahba, 1990] of a) log (*Great Egret*) and b) log (*White Ibis*) as a function of initial water stage and disruption illustrates the nonlinearity of the relationship. Note the generally linear relationship of bird counts with the number of days with disruptions, and the more complex relationship with stage.

**Figure 6** Box plots of the number of birds and the predictors initial water stage (IWS), average water stage (AWS), and disruption (DIS) given three clusters.

**Figure 7** Hierarchical Bayesian Regression Model,  $\mu$  indicates the location parameter (mean) and  $\sigma$  indicates scale parameter (variance).

**Figure 8** Bayesian model predictions (solid line) and observed bird counts at ENP in May (open circle) using AWS, IWS, and DIS values at P33 as predictors for 1985–2006. Model output refers to the posterior mean from the Bayesian model, and Cross-validated value refers to the

posterior mean estimated with the model without using these data (2000-2006).  $r$  values shown at the top right refer to the correlation between the posterior mean and observed values. The uncertainty bounds represent the 5% and the 95% of the posterior conditional distribution.

**Figure 9** Retrospective analysis of wading bird foraging numbers. Model predictions (solid line) and observed birds at ENP in May (open circle), using Bayesian quadratic regression (Eqn. 2) and AWS, IWS, and DIS derived from 1952-2006 water levels at P33. The late 1950s and early 1960s were marked by significant hydrologic manipulations in the regional water management system.

**Figure 10** The long-term trends for P33 average water stage (top), disruption (middle) and rainfall (bottom panel) are indicated by Lowess [Cleveland, 1979] smooth lines superimposed on graphs of time series. The concurrent decadal to multi-decadal variability in the relevant time series is shown.

**Figure 11** Alternative inflow management scenarios produce variable numbers of foraging Great Egret and White Ibis in ENP from 1986 – 2000.

**Table 1** Correlation coefficients for the three predictors initial water stage (IWS), average water stage (AWS) and number of disruptions (DIS) and the two predictands (Great Egret and White Ibis foraging populations), including partial correlations

	Great Egret	White Ibis
<i>Bird, IWS/AWS,DIS</i>	-0.131	0.035
<i>Bird,AWS/IWS,DIS</i>	0.173	-0.136
<i>Bird,DIS/IWS,AWS</i>	0.834	0.861
<i>Bird,IWS/AWS</i>	0.831	0.865
<i>Bird,IWS/DIS</i>	0.117	-0.208
<i>Bird, AWS/IWS</i>	0.831	0.865
<i>Bird,AWS/DIS</i>	0.025	-0.163
<i>Bird, DIS/IWS</i>	0.117	-0.208
<i>Bird,DIS/AWS</i>	0.025	-0.163

† IWS is initial water stage, AWS is averaged water stage and DIS is number of disruptions at P33 from January to April.



**Table 2** Model parameters and associated uncertainty bounds for wading bird foraging populations

	Node	Description	Mean	Stand. Dev.	2.50%	Median	97.50%
<b>Great Egret Population Model</b>	$\beta_1$	<i>Intercept</i>	0.201	0.133	-0.070	0.202	0.461
	$\beta_2$	<i>Initial W.S</i>	-0.018	0.206	-0.415	-0.021	0.374
	$\beta_3$	<i>Initial W.S<sup>2</sup></i>	-0.252	0.077	-0.401	-0.253	-0.099
	$\beta_4$	<i>Averaged W.S</i>	0.382	0.190	0.000	0.388	0.759
	$\beta_5$	<i>Averaged W.S<sup>2</sup></i>	-0.098	0.123	-0.347	-0.099	0.132
	$\beta_6$	<i>No. of Disruption</i>	-0.289	0.087	-0.465	-0.287	-0.122
	$\beta_7$	<i>No. of Disruption<sup>2</sup></i>	-0.102	0.068	-0.238	-0.101	0.029
<b>White Ibis Population Model</b>	$\beta_1$	<i>Intercept</i>	0.146	0.164	-0.194	0.151	0.458
	$\beta_2$	<i>Initial W.S</i>	-0.519	0.265	-1.077	-0.510	-0.021
	$\beta_3$	<i>Initial W.S<sup>2</sup></i>	-0.261	0.087	-0.437	-0.261	-0.094
	$\beta_4$	<i>Averaged W.S</i>	0.649	0.241	0.170	0.643	1.117
	$\beta_5$	<i>Averaged W.S<sup>2</sup></i>	-0.231	0.138	-0.504	-0.231	0.052
	$\beta_6$	<i>No. of Disruption</i>	-0.471	0.111	-0.687	-0.470	-0.260
	$\beta_7$	<i>No. of Disruption<sup>2</sup></i>	-0.094	0.082	-0.252	-0.096	0.073

**Table 3** Regression model performance measures

Statistics	Formula
<b>Correlation Coefficient</b>	$r = \frac{\sum_{t=1}^N (O_t - \bar{O})(P_t - \bar{P})}{\left(\sum_{t=1}^N (O_t - \bar{O})^2\right)^{0.5} \left(\sum_{t=1}^N (P_t - \bar{P})^2\right)^{0.5}}$
<b>Coefficient of Efficiency</b>	$CoE = 1 - \frac{\sum_{t=1}^N (O_t - P_t)^2}{\sum_{t=1}^N (O_t - \bar{O})^2}$
<b>Normalized Bias</b>	$Bias = N^{-1} \sum_{t=1}^N (O_t - P_t) / O_t$

<sup>1</sup>*N* is the sample size and *P* is the model-simulated data and *O* is the observed data.

**Table 4.** Model performance for wading bird foraging predictions.

Predictors	R	CoE	Normalized Bias
<i>Great Egret</i>	0.84	0.72	-2%
<i>White Ibis</i>	0.84	0.71	-6%

**Table 5** Model performance measures, parameters and associated uncertainty bounds for daily predictions of P33 water levels based on hydrologic parameters.

<i>P33 Daily water level prediction</i>	<i>Measure</i>	<i>Cal. 30%</i>	<i>Ver. 70%</i>	<i>Node</i>	<i>Description</i>	<i>Mean</i>	<i>Stand. Dev.</i>	<i>2.50%</i>	<i>Median</i>	<i>97.5%</i>
	<i>r</i>	0.981	0.989	$\beta_1$	<i>Intercept</i>	6.151	0.906	4.823	5.908	7.761
				$\beta_2$	<i>P33<sub>t-1</sub></i>	0.913	0.012	0.892	0.916	0.931
				$\beta_3$	<i>Rainfall<sub>t-1</sub></i>	0.440	0.078	0.289	0.440	0.596
	CoE	0.963	0.978	$\beta_4$	<i>Rainfall<sub>t-1</sub><sup>2</sup></i>	-0.017	0.013	-0.043	-0.017	-0.008
				$\beta_5$	<i>Inflow<sub>t-1</sub></i>	0.398	0.083	0.245	0.396	0.561
	Bias (m)	-0.001	-0.057	$\beta_6$	<i>Inflow<sub>t-1</sub><sup>2</sup></i>	-0.007	0.033	-0.069	-0.007	0.058

**Table 6** Mean and (standard deviation) of predicted hydrologic variables and Great Egret and White Ibis foraging populations for current, baseline conditions and five alternative ENP water management scenarios

1986-1990*					
Scenario	AWS	IWS	DIS	Great Egret	White Ibis
<i>Baseline</i>	-0.93 (0.51)	-1.16 (0.60)	0.52 (0.56)	761 (415)	701 (443)
<i>No inflows</i>	-1.20 (0.39)	-1.40 (0.48)	0.59 (0.38)	517 (199)	441 (200)
<i>Minimum inflows</i>	-0.94 (0.46)	-1.09 (0.51)	0.73 (0.38)	689 (283)	617 (289)
<i>Rainfall plan</i>	-0.95 (0.51)	-1.19 (0.58)	0.31 (0.68)	768 (445)	748 (599)
<i>IOP</i>	-0.85 (0.61)	-1.17 (0.61)	0.31 (0.59)	801 (381)	714 (414)
<i>CSOP</i>	-0.83 (0.62)	-1.19 (0.58)	-0.14 (0.90)	928 (503)	909 (589)

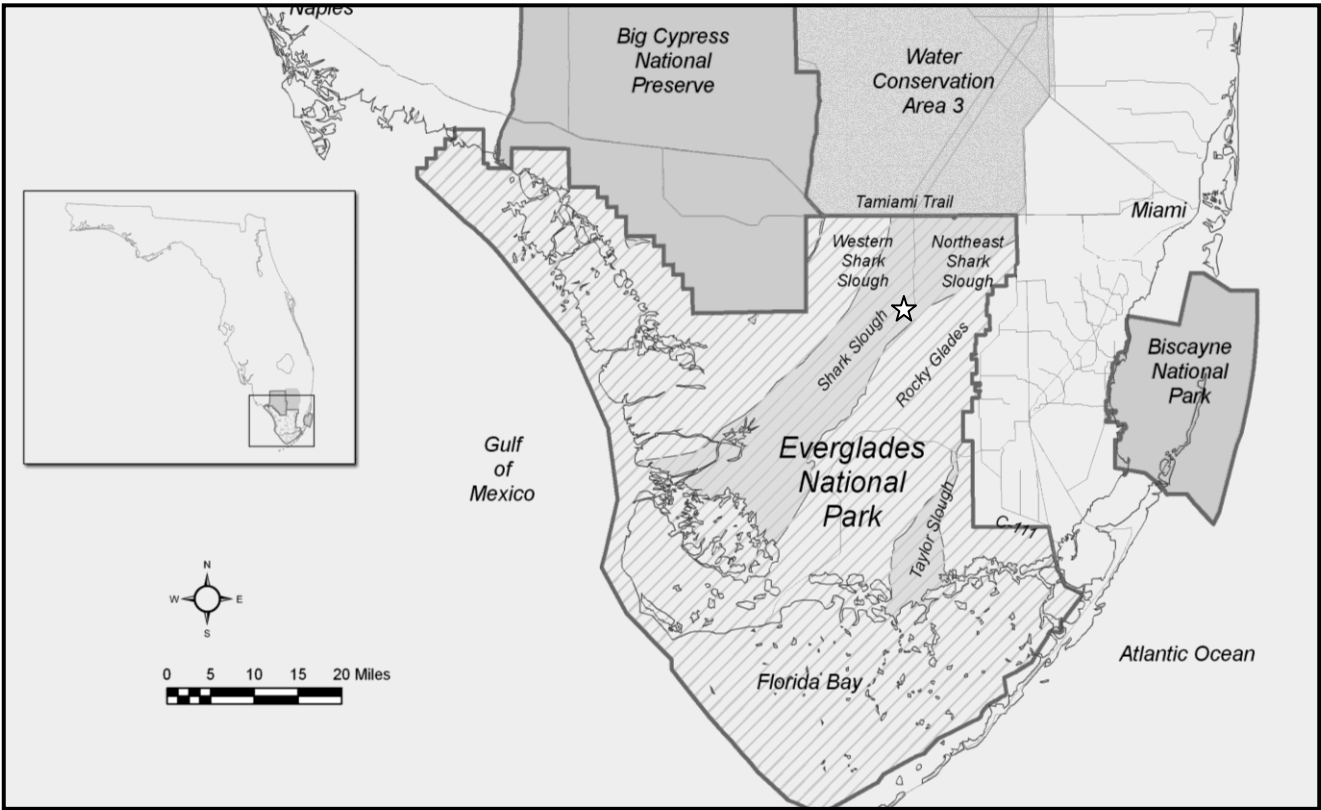
  

1991-2000**					
Scenario	AWS	IWS	DIS	Great Egret	White Ibis
<i>Baseline</i>	0.36 (0.94)	0.20 (1.07)	0.07 (1.10)	1346 (532)	1130 (630)
<i>No inflows</i>	-0.56 (0.45)	-0.56 (0.65)	-0.95 (0.79)	1486 (581)	1898 (914)
<i>Minimum inflows</i>	-0.26 (0.46)	-0.22 (0.66)	-0.76 (0.66)	1774 (570)	2092 (844)
<i>Rainfall plan</i>	0.05 (0.76)	-0.09 (0.87)	-0.72 (0.68)	1698 (607)	1716 (779)
<i>IOP</i>	0.51 (0.99)	0.23 (1.05)	-0.68 (0.80)	1578 (663)	1359 (728)
<i>CSOP</i>	0.51 (0.95)	0.20 (1.11)	-0.49 (0.88)	1488 (563)	1199 (548)

\* Period of below average rainfall

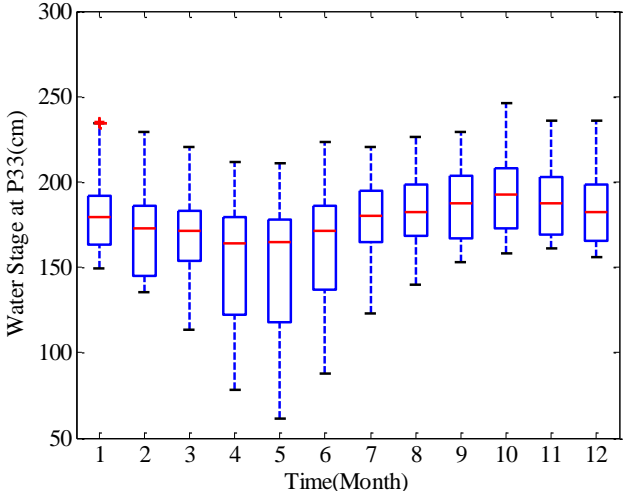
\*\* Period of average to above average rainfall

**Figure 1**

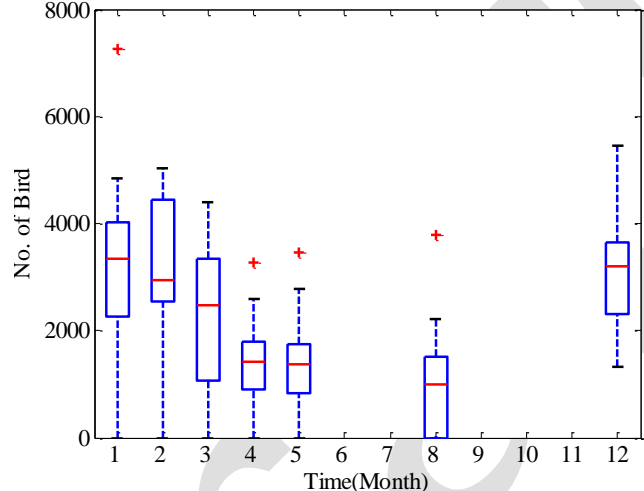


**Figure 2**

*a) Water Stage at P33*



*b) Great Egret*



*c) White Ibis*

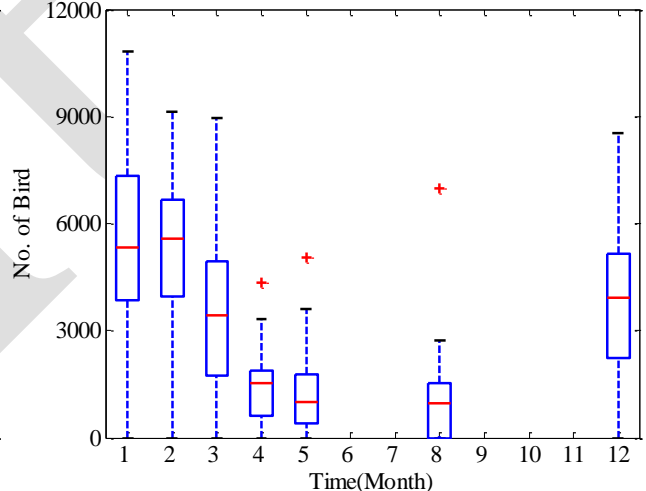
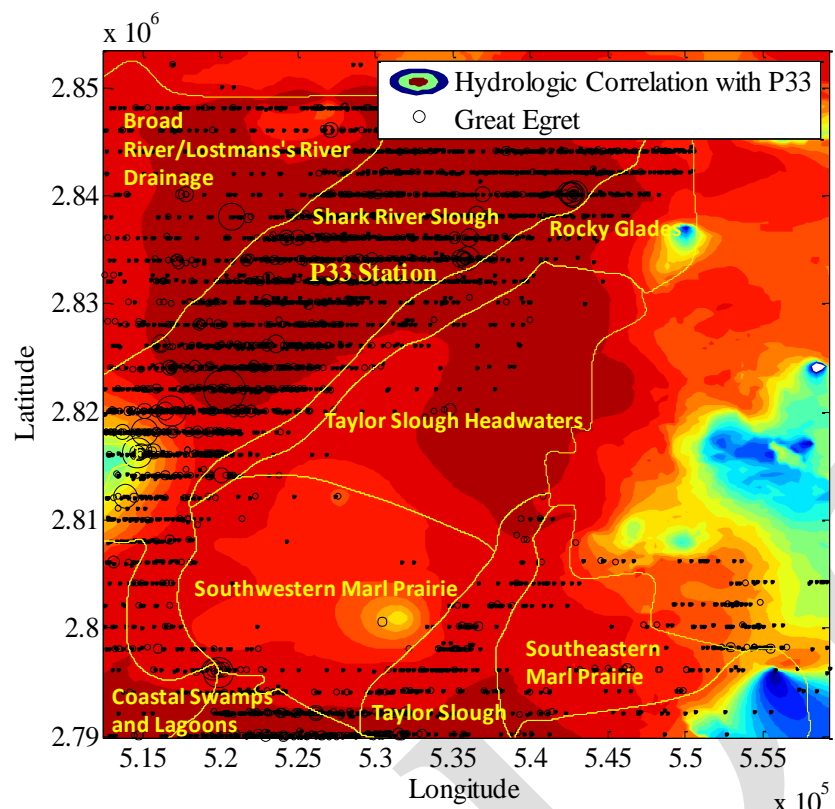
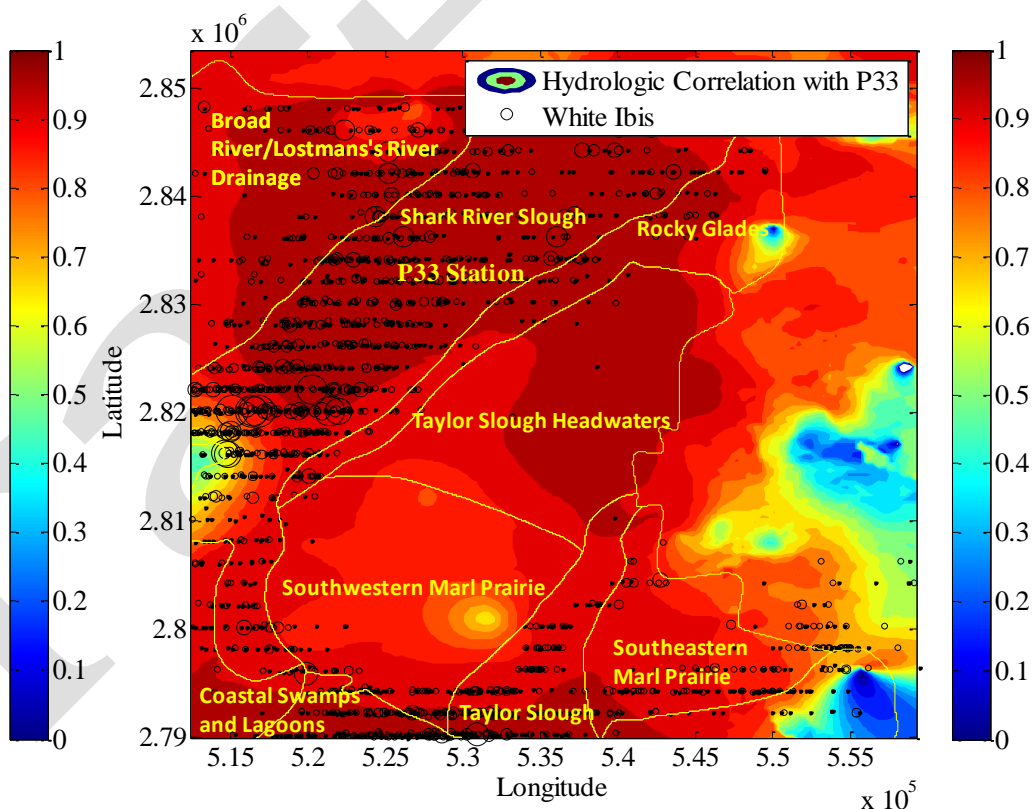


Figure 3

a) Great Egret

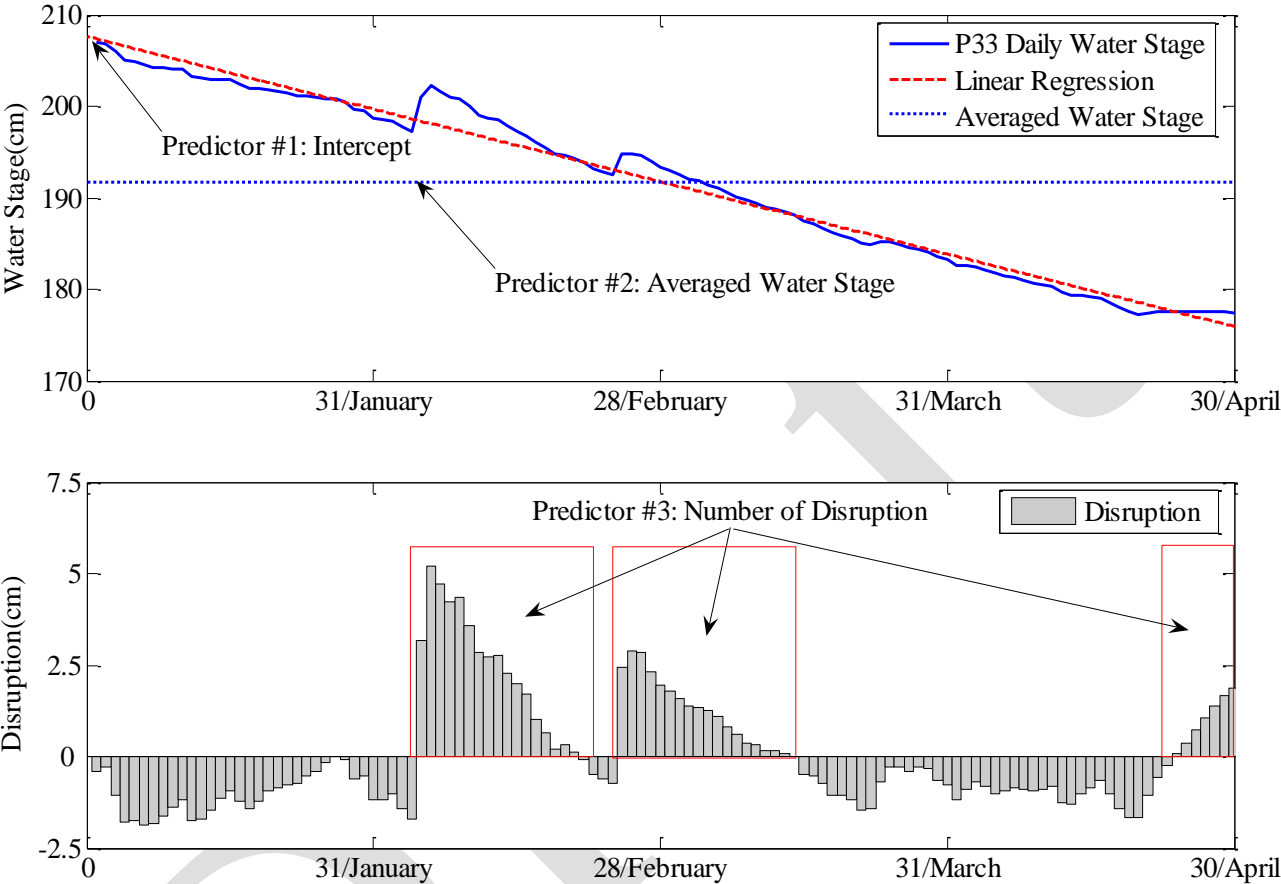


b) White Ibis



831 **Figure 4**

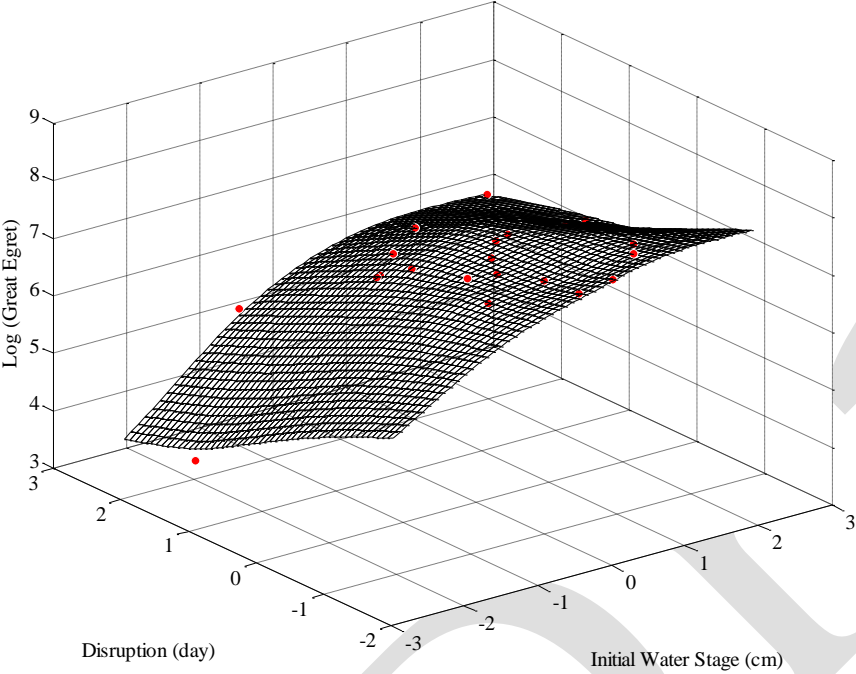
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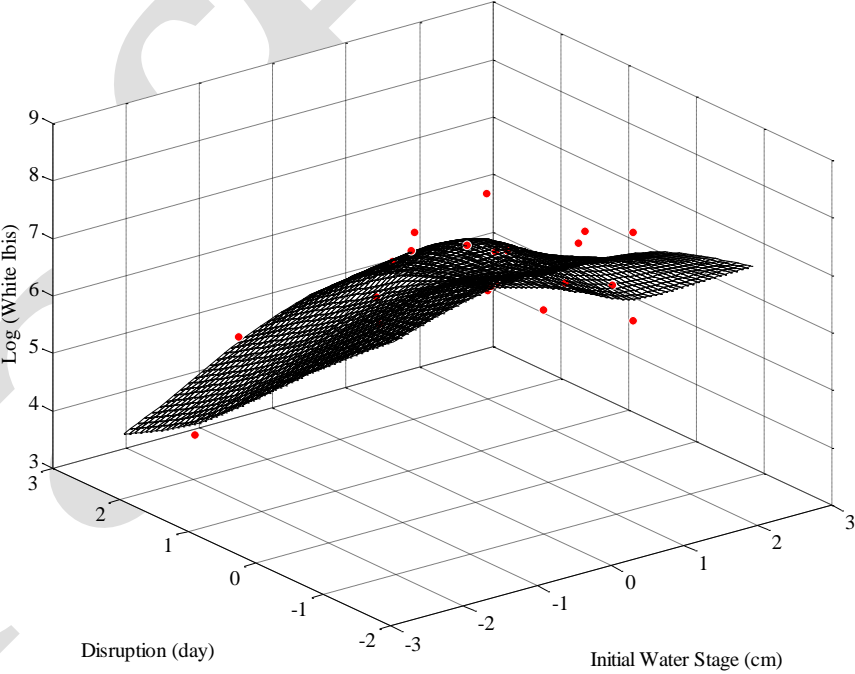
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836 **Figure 5**

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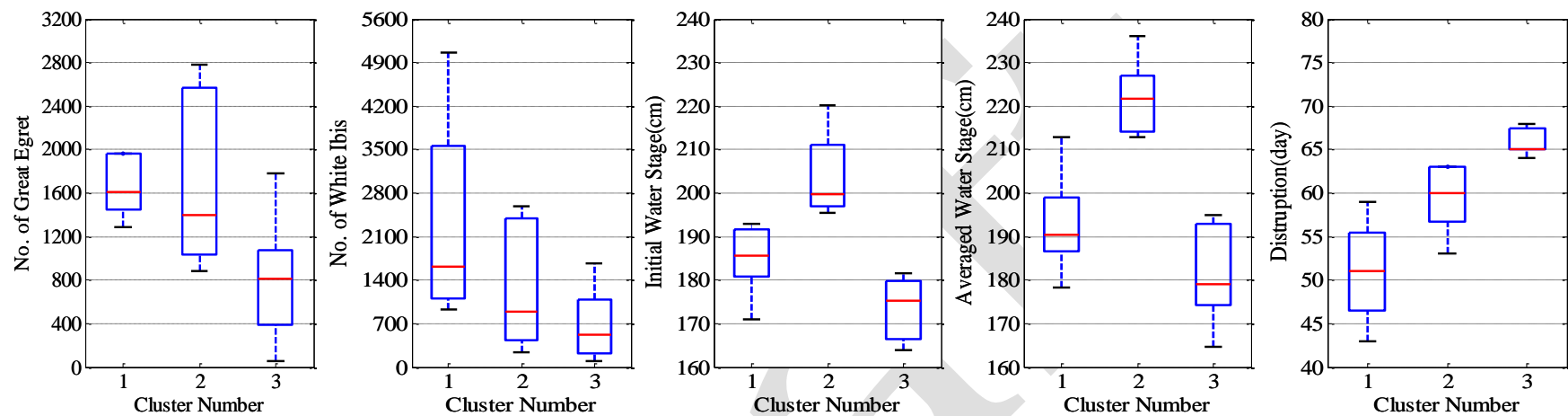
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843 **Figure 6**

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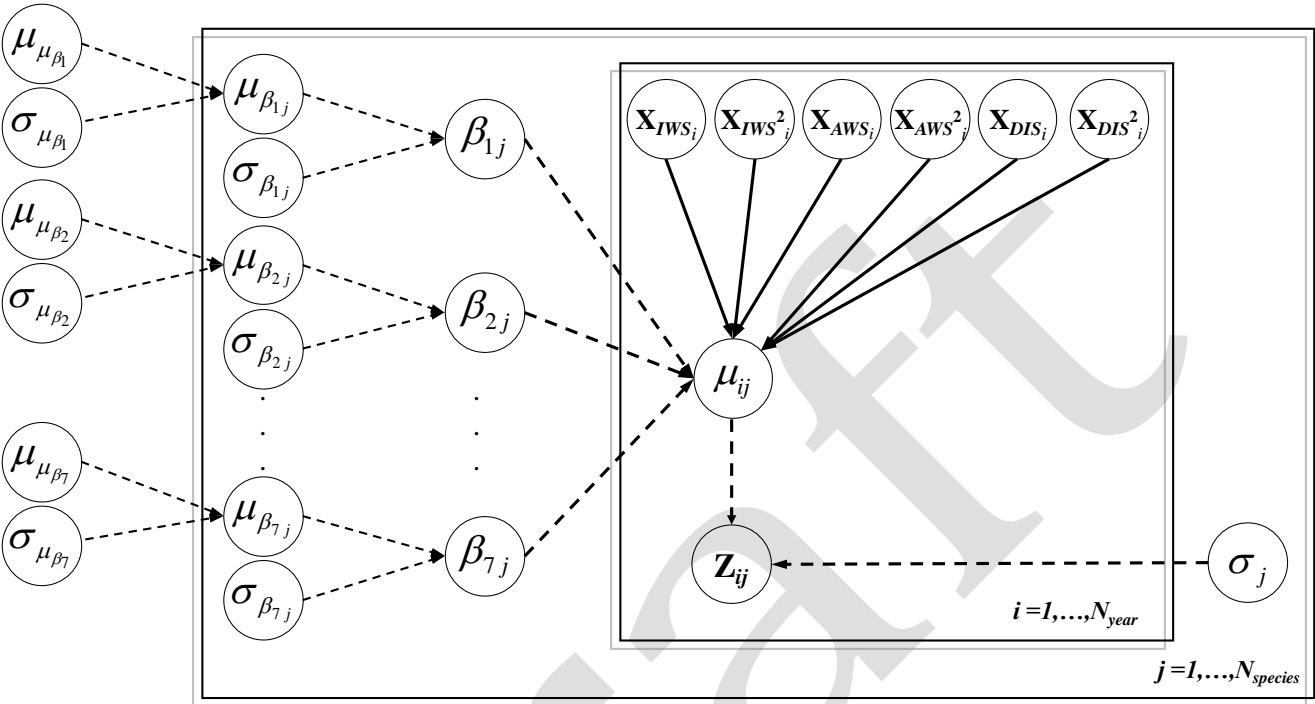


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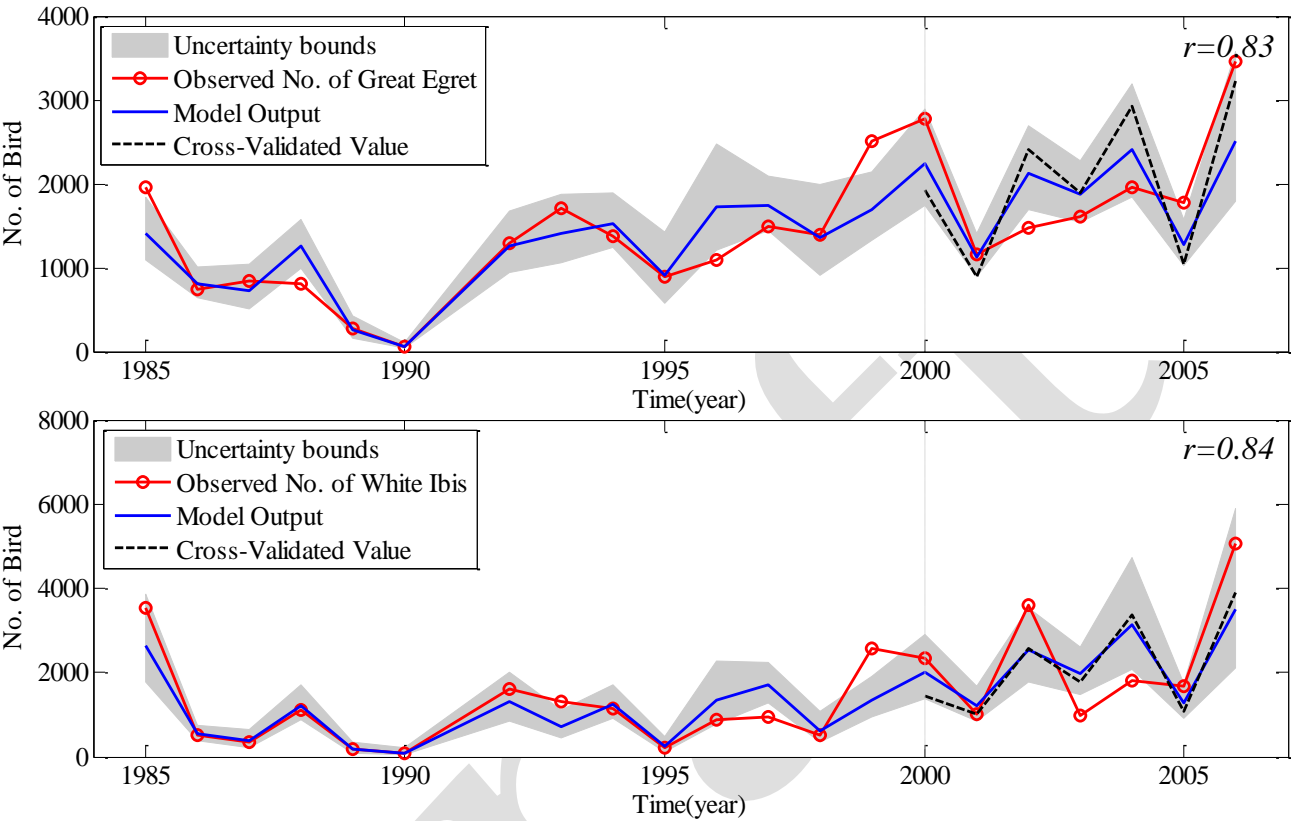


847 **Figure 7**  
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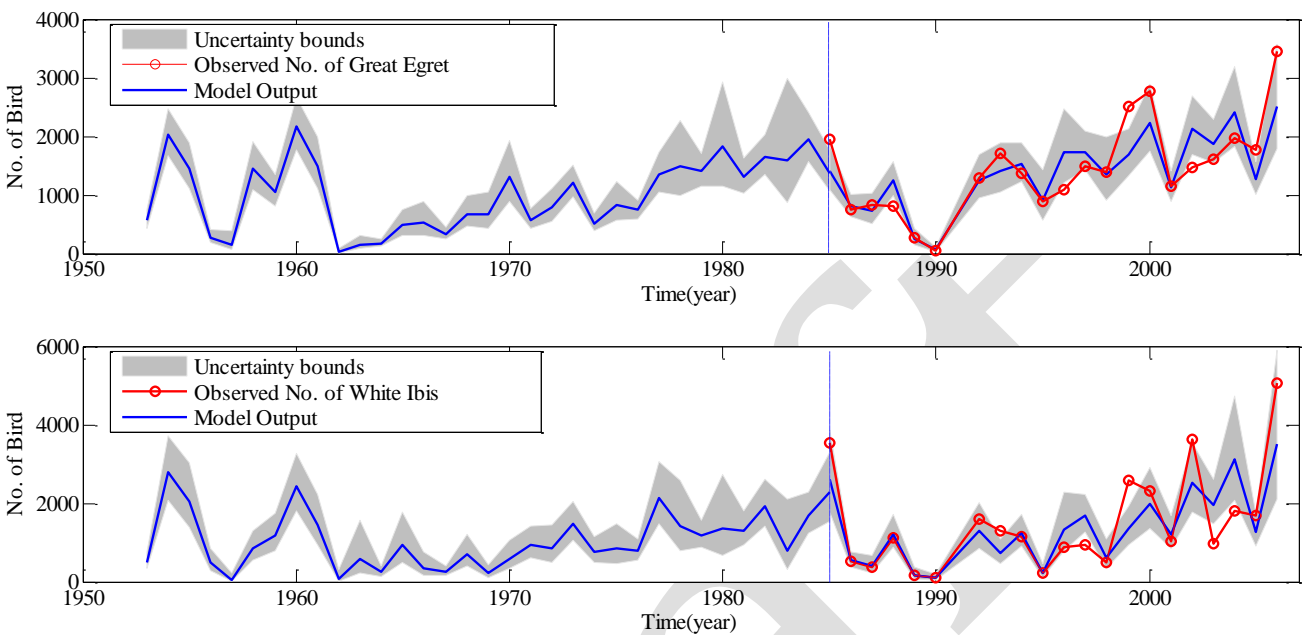
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852 **Figure 8**  
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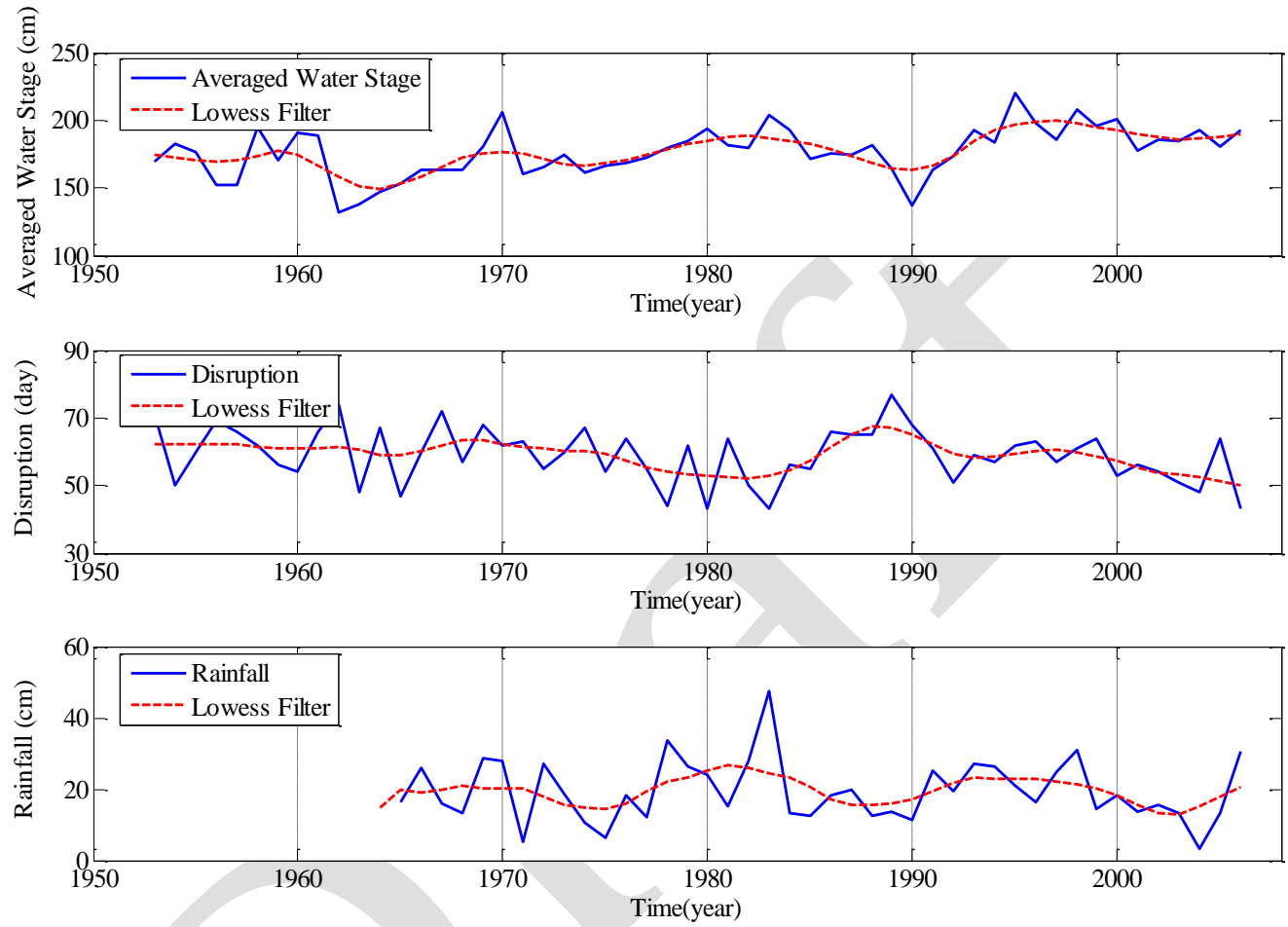
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857 **Figure 9**  
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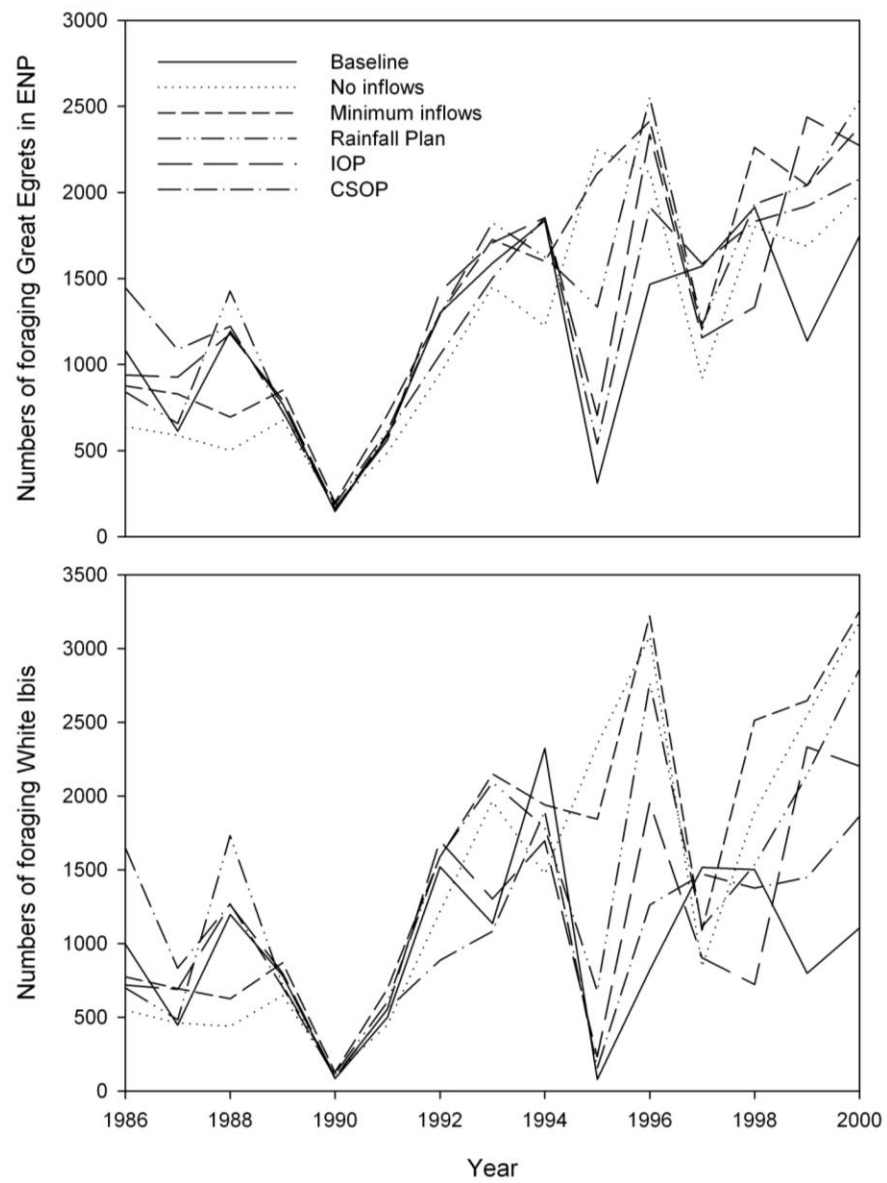
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862 **Figure 10**  
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867 **Figure 11**  
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871 Appendix

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Model hyperparameters and associated uncertainty bounds for each level of the model

Node	Description	Mean	Stand. Dev.	2.50%	Median	97.50%
<b>Hyperparameter Mu Beta (i,1 ) for Great Egret</b>	$\mu_{\beta_{i1}}$ Initial W.S	-0.233	0.428	-1.203	-0.171	0.403
	$\mu_{\beta_{i1}}$ Initial W.S <sup>2</sup>	-0.290	0.253	-0.879	-0.260	0.041
	$\mu_{\beta_{i1}}$ Averaged W.S	0.487	0.548	-0.266	0.516	1.218
	$\mu_{\beta_{i1}}$ Averaged W.S <sup>2</sup>	-0.054	0.580	-0.508	-0.155	2.328
	$\mu_{\beta_{i1}}$ No. of Disruption	-0.362	0.157	-0.717	-0.346	-0.096
	$\mu_{\beta_{i1}}$ No. of Disruption <sup>2</sup>	-0.078	0.135	-0.312	-0.092	0.258
<b>Hyperparameter Mu Beta (i,2 ) for White Ibis</b>	$\mu_{\beta_{i2}}$ Initial W.S	-0.290	0.408	-1.224	-0.265	0.391
	$\mu_{\beta_{i2}}$ Initial W.S <sup>2</sup>	-0.297	0.296	-0.908	-0.262	0.054
	$\mu_{\beta_{i2}}$ Averaged W.S	0.504	0.544	-0.265	0.541	1.220
	$\mu_{\beta_{i2}}$ Averaged W.S <sup>2</sup>	-0.069	0.590	-0.492	-0.179	2.335
	$\mu_{\beta_{i2}}$ No. of Disruption	-0.391	0.165	-0.719	-0.388	-0.088
	$\mu_{\beta_{i2}}$ No. of Disruption <sup>2</sup>	-0.079	0.137	-0.333	-0.093	0.244
<b>Hyperparameter Mu (Mu Beta(i))</b>	$\mu_{\mu_{\beta_i}}$ Initial W.S	-0.262	0.527	-1.237	-0.215	0.417
	$\mu_{\mu_{\beta_i}}$ Initial W.S <sup>2</sup>	-0.291	0.303	-0.916	-0.259	0.090
	$\mu_{\mu_{\beta_i}}$ Averaged W.S	0.497	0.555	-0.280	0.528	1.249
	$\mu_{\mu_{\beta_i}}$ Averaged W.S <sup>2</sup>	-0.068	0.632	-0.512	-0.168	2.341
	$\mu_{\mu_{\beta_i}}$ No. of Disruption	-0.379	0.189	-0.762	-0.369	-0.072
	$\mu_{\mu_{\beta_i}}$ No. of Disruption <sup>2</sup>	-0.076	0.147	-0.350	-0.092	0.266
<b>Hyperparameter Sigma(Mu Beta(i))</b>	$\sigma_{\mu_{\beta_i}}$ Initial W.S	0.001	0.001	0.001	0.002	0.004
	$\sigma_{\mu_{\beta_i}}$ Initial W.S <sup>2</sup>	0.001	0.001	0.000	0.002	0.003
	$\sigma_{\mu_{\beta_i}}$ Averaged W.S	0.001	0.001	0.001	0.002	0.003
	$\sigma_{\mu_{\beta_i}}$ Averaged W.S <sup>2</sup>	0.001	0.001	0.000	0.002	0.003
	$\sigma_{\mu_{\beta_i}}$ No. of Disruption	0.001	0.001	0.000	0.002	0.003
	$\sigma_{\mu_{\beta_i}}$ No. of Disruption <sup>2</sup>	0.001	0.001	0.000	0.002	0.003

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